# Order Determination of Large Dimensional Dynamic Factor Model 

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## Outline

(1) Introduction
(2) Limiting Spectral Distribution
(3) Strong Limit of Extreme Eigenvalues
(4) Application

## Background

Consider the framework of a large dimensional dynamic $k$-factor model with lag $q$

$$
\mathbf{R}_{t}=\sum_{i=0}^{q} \boldsymbol{\Lambda}_{i} \mathbf{F}_{t-i}+\mathbf{e}_{t}, \quad t=1, \ldots, T
$$

- $\boldsymbol{\Lambda}_{i}: n \times k$ non-random matrices with full rank
- $\mathbf{F}_{t}: k \times 1$ iid standard complex random vector
- $\mathbf{e}_{t}: n \times 1$ iid complex, mean zero, variance $\sigma^{2}$, independent of $\mathbf{F}_{t}$
- a information-plus-noise type model
(Dozier \& Silverstein, 2007a, b; Bai \& Silverstein, 2012)
- $n, T \rightarrow \infty$, with $\frac{n}{T} \rightarrow c>0$
- $k, q$ small and fixed but unknown


## Motivation

Under this high dimensional setting, an important statistical problem is to estimate $k$ and $q$ (Bai \& Ng, 2002; Harding, 2012).

## Notations

For fixed $\tau$, define

$$
\begin{aligned}
\boldsymbol{\Phi}_{n}(\tau)= & \frac{1}{2 T} \sum_{j=1}^{T}\left(\mathbf{R}_{j} \mathbf{R}_{j+\tau}^{*}+\mathbf{R}_{j+\tau} \mathbf{R}_{j}^{*}\right) \\
= & \frac{1}{2 T}\left\{\boldsymbol{\Lambda}\left(\mathbf{F}_{0} \mathbf{F}_{\tau}^{\prime}+\mathbf{F}_{\tau} \mathbf{F}_{0}^{\prime}\right) \boldsymbol{\Lambda}^{\prime}\right\}+ \\
& \frac{1}{2 T}\left\{\left(\mathbf{E}_{0} \mathbf{F}_{\tau}^{\prime} \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Lambda} \mathbf{F}_{\tau} \mathbf{E}_{0}^{\prime}\right)+\left(\mathbf{E}_{\tau} \mathbf{F}_{0}^{\prime} \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Lambda} \mathbf{F}_{0} \mathbf{E}_{\tau}^{\prime}\right)\right\}+ \\
& \frac{1}{2 T}\left(\mathbf{E}_{0} \mathbf{E}_{\tau}^{\prime}+\mathbf{E}_{\tau} \mathbf{E}_{0}^{\prime}\right)
\end{aligned}
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& \frac{1}{2 T}\left(\mathbf{E}_{0} \mathbf{E}_{\tau}^{\prime}+\mathbf{E}_{\tau} \mathbf{E}_{0}^{\prime}\right), \\
\mathbf{M}_{n}(\tau)= & \frac{1}{2 T} \sum_{j=1}^{T}\left(\mathbf{e}_{j} \mathbf{e}_{j+\tau}^{*}+\mathbf{e}_{j+\tau} \mathbf{e}_{j}^{*}\right) \\
= & \frac{1}{2 T}\left(\mathbf{E}_{0} \mathbf{E}_{\tau}^{\prime}+\mathbf{E}_{\tau} \mathbf{E}_{0}^{\prime}\right)
\end{aligned}
$$

## Notations

Here,

$$
\begin{aligned}
\boldsymbol{\Lambda} & =\left(\boldsymbol{\Lambda}_{0}, \boldsymbol{\Lambda}_{1}, \cdots, \boldsymbol{\Lambda}_{q}\right)_{n \times k(q+1)}, \\
\mathbf{F}_{\tau} & =\left(\begin{array}{cccc}
\mathbf{F}_{T+\tau} & \mathbf{F}_{T+\tau-1} & \cdots & \mathbf{F}_{\tau+1} \\
\mathbf{F}_{T+\tau-1} & \mathbf{F}_{T+\tau-2} & \cdots & \mathbf{F}_{\tau} \\
\vdots & \vdots & \vdots & \vdots \\
\mathbf{F}_{T+\tau-q} & \mathbf{F}_{T+\tau-1-q} & \cdots & \mathbf{F}_{\tau+1-q}
\end{array}\right)_{k(q+1) \times T}, \\
\mathbf{E}_{\tau} & =\left(\mathbf{e}_{T+\tau}, \mathbf{e}_{T+\tau-1}, \cdots, \mathbf{e}_{\tau+1}\right)_{n \times T} .
\end{aligned}
$$

## Case $\tau=0$

## Fact 1:

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\mathbf{M}_{n}(0)=\frac{1}{T} \sum_{j=1}^{T} \mathbf{e}_{j} \mathbf{e}_{j}^{*}
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- the LSD is MP law (Marčenko and Pastur, 1967) with density

$$
f_{c}(x)=\frac{1}{2 \pi c x} \sqrt{\left(b_{c}-x\right)\left(x-a_{c}\right)}, x \in\left[a_{c}, b_{c}\right]
$$

and a point mass $1-1 / c$ at the origin if $c>1$. Here $c=\lim _{n \rightarrow \infty} \frac{n}{T}, a_{c}=(1-\sqrt{c})^{2}$ and $b_{c}=(1+\sqrt{c})^{2}$.

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## Case $\tau=0$

## Fact 2:

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\text { Recall } \boldsymbol{\Lambda}=\left(\boldsymbol{\Lambda}_{0}, \boldsymbol{\Lambda}_{1}, \cdots, \boldsymbol{\Lambda}_{q}\right)_{n \times k(q+1)}
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\text { Recall } \boldsymbol{\Lambda}=\left(\boldsymbol{\Lambda}_{0}, \boldsymbol{\Lambda}_{1}, \cdots, \boldsymbol{\Lambda}_{q}\right)_{n \times k(q+1)} \\
\Rightarrow \operatorname{Cov} \mathbf{R}_{t}=\sigma^{2} \mathbf{I}+\boldsymbol{\Lambda} \boldsymbol{\Lambda}^{*} \sim\left(\begin{array}{cc}
\sigma^{2} \mathbf{I}+\boldsymbol{\Lambda}^{*} \boldsymbol{\Lambda} & \mathbf{0} \\
\mathbf{0} & \sigma^{2} \mathbf{I}
\end{array}\right)
\end{gathered}
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$$

- a spiked population model (Johnstone, 2001; Baik \& Silverstein, 2006; Bai \& Yao, 2008) with population eigenvalue $\lambda_{1} \geq \lambda_{2} \geq \cdots \geq \lambda_{k(q+1)}>\sigma^{2}=\cdots=\sigma^{2}$.


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- when $\boldsymbol{\Lambda}^{*} \boldsymbol{\Lambda}$ is "not small", sample eigenvalue $\hat{\lambda}_{1} \geq \hat{\lambda}_{2} \geq \cdots \geq \hat{\lambda}_{k(q+1)}>\left(\sigma^{2} b_{c}\right) \geq \hat{\lambda}_{k(q+1)+1} \cdots \geq \hat{\lambda}_{n}$.


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$\Rightarrow$ Can estimate $k(q+1)$ by counting the number of eigenvalues $>\sigma^{2} b_{c}$.


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To do so, we need to investigate the case for at least one $\tau \geq 1$.

## Outline

## (1) Introduction

(2) Limiting Spectral Distribution

## 3 Strong Limit of Extreme Eigenvalues

(4) Application

## Main Result

## Theorem 1 (Jin et al. (2014))

## Assume:

- (a) $\tau \geq 1$ is a fixed integer.
- (b) $\mathbf{e}_{k}=\left(\varepsilon_{1 k}, \cdots, \varepsilon_{n k}\right)^{\prime}, k=1,2, \ldots, T+\tau$, are $n$ dimensional vectors of independent standard complex components with $\sup _{1 \leq i \leq n, 1 \leq t \leq T+\tau} \mathrm{E}\left|\varepsilon_{i t}\right|^{2+\delta} \leq M<\infty$ for some $\delta \in(0,2]$, and for any $\eta>0$,

$$
\begin{equation*}
\frac{1}{\eta^{2+\delta} n T} \sum_{i=1}^{n} \sum_{t=1}^{T+\tau} \mathrm{E}\left(\left|\varepsilon_{i t}\right|^{2+\delta} I\left(\left|\varepsilon_{i t}\right| \geq \eta T^{1 /(2+\delta)}\right)\right)=o(1) \tag{1}
\end{equation*}
$$

- (c) $n /(T+\tau) \rightarrow c>0$ as $n, T \rightarrow \infty$.
- (d) $\mathbf{M}_{n}=\sum_{k=1}^{T}\left(\gamma_{k} \gamma_{k+\tau}^{*}+\gamma_{k+\tau} \gamma_{k}^{*}\right)$, where $\gamma_{k}=\frac{1}{\sqrt{2 T}} \mathbf{e}_{k}$.


## Main Result

## Theorem 1 (Jin et al. (2014)) (cont'd)

Then as $n, T \rightarrow \infty, F^{\mathrm{M}_{n}} \xrightarrow{D} F_{\tau}$ a.s. and $F_{\tau}$ has a density function

$$
\phi_{c}(x)=\frac{1}{2 c \pi} \sqrt{\frac{y_{0}^{2}}{1+y_{0}}-\left(\frac{1-c}{x}+\frac{1}{\sqrt{1+y_{0}}}\right)^{2}},|x| \leq d_{c}
$$

where

$$
d_{c}=\left\{\begin{array}{cl}
\frac{(1-c) \sqrt{1+y_{1}}}{y_{1}-1}, & c \neq 1 \\
2, & c=1
\end{array}\right.
$$

$y_{0}$ is the largest real root of the equation:
$y^{3}-\frac{(1-c)^{2}-x^{2}}{x^{2}} y^{2}-\frac{4}{x^{2}} y-\frac{4}{x^{2}}=0$;
and $y_{1}$ is the only real root of the equation:
$\left((1-c)^{2}-1\right) y^{3}+y^{2}+y-1=0$
such that $y_{1}>1$ if $c<1$ and $y_{1} \in(0,1)$ if $c>1$.
Further, if $c>1$, then $F_{\tau}$ has a point mass $1-1 / c$ at the origin.

## Main Result



Figure 1: $\phi_{c}(x)$ with $c=0.2$ (black), 0.5 (blue) and 0.7 (red).


Figure 2 : $\phi_{c}(x)$ with $c=1.5$ (black), 2 (blue) and 2.5 (red). The area under each curve is $1 / c$.

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## Motivation

Once the LSD of $\mathbf{M}_{n}(\tau)$ is derived, it is observed that the number of eigenvalues of $\boldsymbol{\Phi}_{n}(\tau)$ that lie outside the support of the LSD of $\mathbf{M}_{n}(\tau)$ at lags $1 \leq \tau \leq q$ is different from that at lags $\tau>q$. Thus, the estimates of $k$ and $q$ can be separated by counting the number of eigenvalues of $\boldsymbol{\Phi}_{n}(\tau)$ that lie outside the support of the LSD of $\mathbf{M}_{n}(\tau)$ from $\tau=0,1,2, \cdots, q$, $q+1, \cdots$.

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It is worth noting that for this method to work, we require that with probability 1 , there is no eigenvalues outside the the support of the LSD of $\mathbf{M}_{n}(\tau)$ so that if an eigenvalue of $\boldsymbol{\Phi}_{n}(\tau)$ goes out of the support of the LSD of $\mathbf{M}_{n}(\tau)$, it must come from the information part.

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This motivates us to establish the limits of the largest and smallest eigenvalues of $\mathbf{M}_{n}(\tau)$, after showing that with probability 1 no eigenvalues exist outside the support of the LSD of $\mathbf{M}_{n}(\tau)$.

## Main Results

## Theorem 2

Assume:

- (a) $\tau \geq 1$ is a fixed integer.
- (b) $\mathbf{e}_{k}=\left(\varepsilon_{1 k}, \cdots, \varepsilon_{n k}\right)^{\prime}, k=1,2, \ldots, T+\tau$, are $n$-vectors of independent standard complex components with $\sup _{i, t} \mathrm{E}\left|\varepsilon_{i t}\right|^{4} \leq M$ for some $M>0$.
- (c) There exist $K>0$ and a random variable $X$ with finite fourth order moment such that, for any $x>0$, for all $n, T$ $\frac{1}{n T} \sum_{i=1}^{n} \sum_{t=1}^{T+\tau} \mathrm{P}\left(\left|\varepsilon_{i t}\right|>x\right) \leq K \mathrm{P}(|X|>x)$.
- (d) $c_{n} \equiv n / T \rightarrow c>0$ as $n \rightarrow \infty$.
- (e) $\mathbf{M}_{n}=\sum_{k=1}^{T}\left(\gamma_{k} \gamma_{k+\tau}^{*}+\gamma_{k+\tau} \gamma_{k}^{*}\right)$, where $\gamma_{k}=\frac{1}{\sqrt{2 T}} \mathbf{e}_{k}$.
- (f) The interval $[\mathrm{a}, \mathrm{b}]$ lies outside the support of $F_{\tau}$.

Then $\mathrm{P}\left(\right.$ no eigenvalues of $\mathbf{M}_{n}$ appear in $[a, b]$ for all large n$)=1$.

## Main Results

## Theorem 3

Assuming conditions (a)-(e) in Theorem 2 hold, we have

$$
\lim _{n \rightarrow \infty} \lambda_{\min }\left(\mathbf{M}_{n}\right)=-d_{c} \quad \text { a.s. } \quad \text { and } \quad \lim _{n \rightarrow \infty} \lambda_{\max }\left(\mathbf{M}_{n}\right)=d_{c} \quad \text { a.s. }
$$

Here, $-d_{c}$ and $d_{c}$ are the left and right boundary points of the support of the LSD of $\mathbf{M}_{n}$, as defined in Theorem 1.

## Simulation



Figure 3: $\phi_{c}(x)$ and plot of sample eigenvalues with $\tau=1, c=0.2$ ( $n=200, T=1000$ ).


Figure 4: $\phi_{c}(x)$ and plot of sample eigenvalues with $\tau=1, c=2.5$

$$
(n=2500, T=1000) .
$$

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## Estimation of $k$ and $q$

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Recall that $k(q+1)$ can be estimated by counting the number of spiked eigenvalues of $\boldsymbol{\Phi}_{n}(0)$.

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For $\tau \geq 1$, we have

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\begin{aligned}
\boldsymbol{\Phi}_{n}(\tau)= & \frac{1}{2 T}\left\{\boldsymbol{\Lambda}\left(\mathbf{F}_{0} \mathbf{F}_{\tau}^{\prime}+\mathbf{F}_{\tau} \mathbf{F}_{0}^{\prime}\right) \boldsymbol{\Lambda}^{\prime}\right\}+ \\
& \frac{1}{2 T}\left\{\left(\mathbf{E}_{0} \mathbf{F}_{\tau}^{\prime} \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Lambda} \mathbf{F}_{\tau} \mathbf{E}_{0}^{\prime}\right)+\left(\mathbf{E}_{\tau} \mathbf{F}_{0}^{\prime} \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Lambda} \mathbf{F}_{0} \mathbf{E}_{\tau}^{\prime}\right)\right\}+ \\
& \mathbf{M}_{n}
\end{aligned}
$$

## Estimation of $k$ and $q$

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Define $\mathbf{B}_{1}=\mathbf{\Lambda Q}$ and $\mathbf{B}=\left(\mathbf{B}_{1} \vdots \mathbf{B}_{2}\right)$ is an $n \times n$ orthogonal matrix, where $\mathbf{Q}=\left(\boldsymbol{\Lambda}^{\prime} \boldsymbol{\Lambda}\right)^{-1 / 2}$.

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Then, $\mathbf{B}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}=\left(\begin{array}{ll}\mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} & \mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2} \\ \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} & \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}\end{array}\right)$.

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Then, $\mathbf{B}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}=\left(\begin{array}{ll}\mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} & \mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2} \\ \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} & \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}\end{array}\right)$.
Note that $\mathbf{B}_{2} \boldsymbol{\Lambda}=0$, we have

$$
\begin{aligned}
& \mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} \sim \mathbf{Q A}_{\tau} \mathbf{Q}+\mathbf{B}_{1}^{\prime} \mathbf{M}_{n} \mathbf{B}_{1} \\
& \mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}=\mathbf{B}_{1}^{\prime} \mathbf{M}_{n} \mathbf{B}_{2}+\frac{1}{2 T} \mathbf{Q}\left(\mathbf{F}_{0} \mathbf{E}_{\tau}^{*}+\mathbf{F}_{\tau} \mathbf{E}_{0}^{*}\right) \mathbf{B}_{2} \\
& \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1}=\mathbf{B}_{2}^{\prime} \mathbf{M}_{n} \mathbf{B}_{1}+\frac{1}{2 T} \mathbf{B}_{2}^{\prime}\left(\mathbf{E}_{0} \mathbf{F}_{\tau}^{*}+\mathbf{E}_{\tau} \mathbf{F}_{0}^{*}\right) \mathbf{Q} \\
& \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}=\mathbf{B}_{2}^{\prime} \mathbf{M}_{n} \mathbf{B}_{2} .
\end{aligned}
$$

where $\left(\mathbf{A}_{\tau}\right)_{k(q+1) \times k(q+1)}$ is the matrix with 1's on upper and lower $k \tau$ diagonals and 0's elsewhere.

## Estimation of $k$ and $q$

If $\ell$ is a root of $\boldsymbol{\Phi}_{n}(\tau)$ but not a root of $\mathbf{B}_{2}^{\prime} \mathbf{M}_{n} \mathbf{B}_{2}$, then

$$
0=\left|\begin{array}{cc}
\mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1}-\ell \mathbf{l} & \mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2} \\
\mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1} & \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}-\ell \mathbf{I}
\end{array}\right|
$$

Since $\left|\mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}-\ell \mathbf{I}\right|=\left|\mathbf{B}_{2}^{\prime} \mathbf{M}_{n} \mathbf{B}_{2}-\ell \mathbf{I}\right| \neq 0$, we have

$$
\left|\mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1}-\ell \mathbf{I}-\mathbf{B}_{1}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{2}\left(\mathbf{B}_{2}^{\prime} \mathbf{M}_{n} \mathbf{B}_{2}-\ell \mathbf{I}\right)^{-1} \mathbf{B}_{2}^{\prime} \boldsymbol{\Phi}_{n}(\tau) \mathbf{B}_{1}\right|=0
$$

After certain simplification, the equation above can be shown equivalent to

$$
\left|\mathbf{A}_{\tau}-\left(\ell+\frac{c m(\ell)}{1-c^{2} m^{2}(\ell)+\sqrt{\left.1-c^{2} m^{2}(\ell)\right)}}\right) \mathbf{Q}^{-2}-\frac{c m(\ell)}{2 \sqrt{1-c^{2} m^{2}(\ell)}} \mathbf{I}_{k \times(q+1)}\right|
$$

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The above equation is the key relation between signals and the observed spikes.

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However, if the matrices $\mathbf{A}_{\tau}$ and $\mathbf{Q}^{2}$ are commutative, the transition phenomenon becomes clear, that is, there is a common orthogonal matrix $\mathbf{O}$ to simultaneously diagonalize the two matrices, i.e., we have $\mathbf{A}_{\tau}=\mathbf{O} \mathbf{D}_{\tau} \mathbf{O}^{\prime}$ and $\mathbf{Q}^{2}=\mathbf{O} \mathbf{D}_{\lambda} \mathbf{O}^{\prime}$, where $\mathbf{D}_{\tau}=\operatorname{diag}\left[a_{1}, \cdots, a_{k(q+1)}\right]$ and $\mathbf{D}_{\lambda}=\operatorname{diag}\left[\lambda_{1}, \cdots, \lambda_{k(q+1)}\right]$.

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Then, the equation becomes

$$
\begin{gathered}
a_{j}-\left(\ell+\frac{c m(\ell)}{1-c^{2} m^{2}(\ell)+\sqrt{\left.1-c^{2} m^{2}(\ell)\right)}}\right) \lambda_{j}^{-1}-\frac{c m(\ell)}{2 \sqrt{1-c^{2} m^{2}(\ell)}}=0, \\
j=1,2 \cdots, k(q+1) .
\end{gathered}
$$

## Estimation of $k$ and $q$

Case 1. If $a_{j} \geq 0$ and $g(d(c))>a_{j}$, then the equation $a_{j}=g(\ell)$ doesn't have a solution in the interval $(d(c), \infty)$ because $g(\ell)$ is increasing and continuous, where

$$
g(\ell)=\left(\ell+\frac{c m(\ell)}{1-c^{2} m^{2}(\ell)+\sqrt{\left.1-c^{2} m^{2}(\ell)\right)}}\right) \lambda_{j}^{-1}+\frac{c m(\ell)}{2 \sqrt{1-c^{2} m^{2}(\ell)}} .
$$

On the interval $(-\infty,-d(c))$ it does not have solution either because $g(\ell)<g(-d(c))=-g(d(c))<0$. Thus, the equation $a_{j}=g(\ell)$ does not have any solution.

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Case 2. If $a_{j} \geq 0$ and $a_{j} \geq g(d(c))>0$, then the equation $a_{j}=g(\ell)$ has a solution in the interval $(d(c), \infty)$, and on the interval $(-\infty,-d(c))$ it does not have solution either because $a_{j} \geq 0$ and $g(\ell) \leq g(-d(c))<0$. Thus, the equation $a_{j}=g(\ell)$ has only one solution.

## Estimation of $k$ and $q$

Case 3. If $a_{j} \geq 0$ and $a_{j}>-g(d(c)) \geq 0$, then the equation $a_{j}=g(\ell)$ has a solution in the interval $(d(c), \infty)$, and on the interval $(-\infty,-d(c))$ it does not have any solution because $a_{j}>g(-d(c)) \geq g(\ell)$ when $\ell<-d(c)$.

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Case 4. If $-g(d(c)) \geq a_{j} \geq g(d(c))$, then the equation $a_{j}=g(\ell)$ has a solution in the interval $(d(c), \infty)$ and another solution on the interval $(-\infty,-d(c))$. Especially when $a_{j}=0$, the case is true.

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Similarly, we may discuss the cases when $a_{j} \leq 0$.
Since $m(d(c))<0$, we have $g(d(c))<0$ provided that $\lambda_{j}$ is large enough. Thus, case 1 doesn't happen in general.

## Estimation of $k$ and $q$

- Therefore, the number of spiked eigenvalues of $\boldsymbol{\Phi}_{n}(\tau)$ satisfies

$$
\hat{p}(\tau) \rightarrow\left\{\begin{array}{l}
k(q+1), \quad \tau=0 \\
2 k(q+1)-h(\tau), \quad 1 \leq \tau \leq q \\
2 k(q+1), \quad \tau>q
\end{array}\right.
$$

where $h(\tau)=2 . \#\left\{j, g(d(c))>\left|a_{j}\right|\right\}+{ }^{\#}\left\{j,\left|a_{j}\right|>|g(d(c))|>0\right\}$.

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- Generally, the first case doesn't happen, unless $\lambda_{j}$ is very small.
- Transition threshold:

$$
\lambda_{0}(c)=-\frac{2 \sqrt{1-c^{2} m^{2}(d(c))}\left(d(c)+\frac{c m(d(c))}{\left(1-c^{2} m^{2}(d(c))\right)+\sqrt{1-c^{2} m^{2}(d(c))}}\right)}{c m(d(c))}
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- That is, when $\lambda_{j}>\lambda_{0}(c)$, then $g_{j}(d(c))<0$.


## Estimation of $k$ and $q$

## Algorithm

- Count the number of spiked eigenvalues of $\boldsymbol{\Phi}_{n}(0), \widehat{k(q+1)}$.
- For $\tau=1,2, \cdots$, count the number of spiked eigenvalues of $\boldsymbol{\Phi}_{n}(\tau)$ and stop at the smallest lag $\widehat{q+1}$, at which the number jumps to $2 k(\overline{q+1})$.
- Set $\hat{k}=\frac{k\left(\frac{q+1)}{q+1}\right.}{}$ and $\hat{q}=\widehat{q+1}-1$.


## Simulation



Figure 5: Sample eigenvalues plots for a factor model with no factors with $n=450, T=500, k=0, q=0$ and $\sigma_{\varepsilon}^{2}=1$.

## Simulation



Figure 6: Sample eigenvalues plots for a factor model with $n=450, T=500$, $k=2, q=0$ and $\sigma_{\varepsilon}^{2}=1$.

## Simulation



Figure 7: Sample eigenvalues plots for a factor model with $n=450, T=500$, $k=2, q=1$ and $\sigma_{\varepsilon}^{2}=1$.

## Simulation

| $\tau=0$ | $\tau=1$ | $\tau=2$ | $\tau=3$ | $\tau=4$ | $\tau=5$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 10.4834 | 13.2847 | 9.6398 | 4.5707 | 4.1978 | 4.3652 |
| 10.1067 | 12.8983 | 9.3054 | 4.5100 | 3.7982 | 4.2543 |
| 9.5428 | 10.5731 | 8.9893 | 3.7854 | 3.5023 | 3.6568 |
| 8.1918 | 9.7048 | 8.8115 | 3.4964 | 3.2956 | 3.2796 |
| 7.8733 | 2.7132 | 3.0196 | 3.4424 | 2.9948 | 3.2131 |
| 7.6733 | 2.2934 | 2.8472 | 3.2752 | 2.8658 | 3.0014 |
| 1.8057 | 2.0844 | 2.7571 | 3.1088 | 2.8206 | 2.9009 |
| 1.7851 | 1.9410 | 2.7238 | 2.4418 | 2.6166 | 2.7364 |
| 1.7475 | 1.7971 | 1.8099 | 2.4222 | 2.6032 | 2.5338 |
| 1.7273 | 1.7096 | 1.7313 | 2.3283 | 2.4414 | 2.1618 |
| 1.7090 | 1.7068 | 1.7232 | 2.1798 | 2.3751 | 2.1310 |
| 1.6787 | 1.6803 | 1.6998 | 2.0149 | 2.1294 | 1.9938 |
| 1.6619 | 1.6418 | 1.6874 | 1.8028 | 1.7561 | 1.7109 |

Table 1: Absolute values of the largest eigenvalues of $\boldsymbol{\Phi}_{n}$ at various lags, for $c=0.9, b_{c}=(1+\sqrt{c})^{2}=3.7974, d_{c}=1.8573$.

## Thank you!

