

Mesoscopic eigenvalue correlations of random matrices

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Eigenvalue correlation functions

Hermitian $N \times N$ random matrix H with eigenvalues $\lambda_1, \dots, \lambda_N$.

Goal: eigenvalue density-density correlations at two different energies.



Eigenvalue process $\sum_i \delta_{\lambda_i}$, with **correlation functions**

$$\rho_1(x) := \mathbb{E} \sum_i \delta(x - \lambda_i), \quad \rho_2(x, y) := \mathbb{E} \sum_{i \neq j} \delta(x - \lambda_i) \delta(y - \lambda_j), \quad \dots$$

Connected (or truncated) two-point function

$$p(x, y) := \rho_2(x, y) - \rho_1(x)\rho_1(y),$$

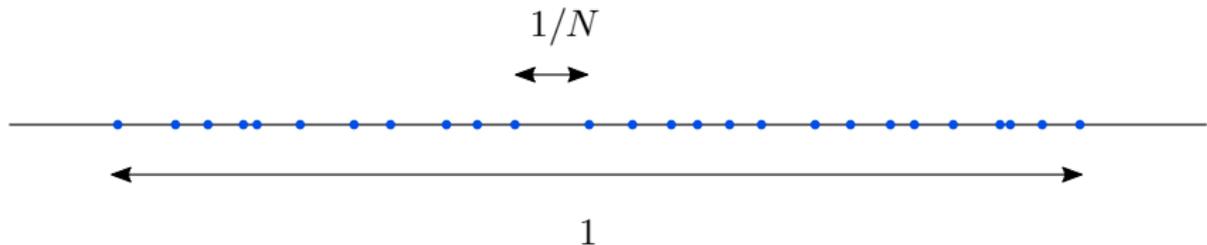
measures eigenvalue density-density correlations at the energies x and y .

Spectral scales

Let H be a **Wigner matrix**: $(H_{ij} : i \leq j)$ are independent with $\mathbb{E}H_{ij} = 0$, $\mathbb{E}|\sqrt{N}H_{ij}|^2 = 1$, and $|\sqrt{N}H_{ij}|^C = O(1)$.

Scales $\omega = \frac{y-x}{2}$:

- **Macroscopic**: $\omega \asymp 1$ (global extent of the spectrum).
- **Microscopic**: $\omega \asymp 1/N$ (eigenvalue spacing).
- **Mesoscopic**: $1/N \ll \omega \ll 1$.



Microscopic eigenvalue correlations

Let $\varrho_E := \frac{1}{2\pi} \sqrt{(4 - E^2)_+}$ be the semicircle law.

To analyse the **microscopic correlations**, choose an energy E and consider **rescaled eigenvalue process**

$$\sum_i \delta_{N\varrho_E(\lambda_i - E)}$$

with connected two-point function

$$p_E(u, v) = \frac{1}{(N\varrho_E)^2} p\left(E + \frac{u}{N\varrho_E}, E + \frac{v}{N\varrho_E}\right).$$

Microscopic density-density correlations \longleftrightarrow behaviour of p_E for fixed u, v .

Wigner-Gaudin-Mehta-Dyson (WGDM) statistics

Let H be GUE ($\beta = 2$) or GOE ($\beta = 1$). Then

$$\lim_{N \rightarrow \infty} p_E(u, v) = Y_\beta(u - v) \quad (1)$$

weakly, where

$$Y_2(u) := -s(u)^2, \quad Y_1(u) := -s'(u) \int_u^\infty s(v) dv - s(u)^2,$$

with the **sine kernel**

$$s(u) := \frac{\sin(\pi u)}{\pi u}.$$

For $u \rightarrow \infty$ we have the asymptotic expansion

$$Y_1(u) = -\frac{1}{\pi^2 u^2} + \frac{1 + \cos^2(\pi u)}{\pi^4 u^4} + O\left(\frac{1}{u^6}\right).$$

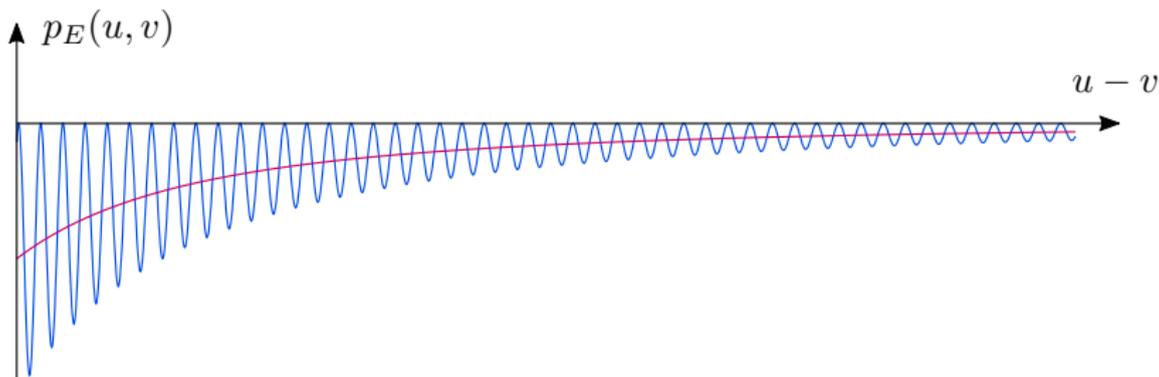
Theorem [Universality (ABEKSTVYY, 2009-2016)]. (1) holds for arbitrary Wigner matrices.

WGMD statistics for mesoscopic separations

Goal:

WGMD statistics for p_E on mesoscopic scales $1 \ll u - v \leq N$?

I.e., analyse the density-density correlations between energies u and v at mesoscopic separations.



Macroscopic fluctuations of linear statistics

A different type of fluctuation result, involving correlations of $O(N^2)$ eigenvalues of typical macroscopic separation $\omega \asymp 1$.

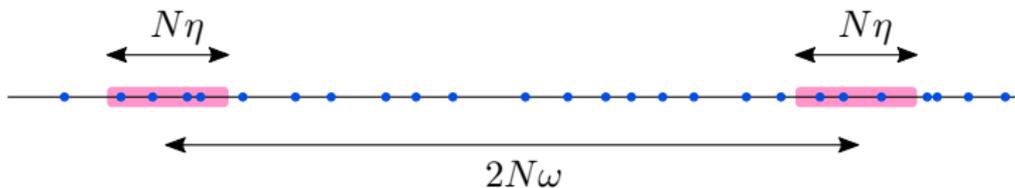
E.g. for $\beta = 1$ we have [Lytova, Pastur – 2009]

$$\text{Var} \sum_i f(\lambda_i) = \frac{1}{2\pi^2} \int_{-2}^2 dx \int_{-2}^2 dy \left(\frac{f(x) - f(y)}{x - y} \right)^2 \frac{4 - xy}{\sqrt{4 - x^2} \sqrt{4 - y^2}} + \frac{N^2 \mathcal{C}_4(H_{12})}{2\pi^2} \left(\int_{-2}^2 dx f(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \right)^2,$$

where $\mathcal{C}_4(\cdot)$ is the fourth cumulant.

Any relation to WGDM statistics?

Result



Fix $E \in (-2, 2)$ and $f, g \in C_c^\infty(\mathbb{R})$. Let

$$1/N \ll \eta \leq \omega \leq 1.$$

Using the notation

$$f_\pm(u) := \frac{1}{N\eta} f\left(\frac{u \mp N\omega}{N\eta}\right).$$

we have

$$\int p_E(u, v) f_+(u) g_-(v) du dv = \int \Upsilon_{E, \beta}(u, v) f_+(u) g_-(v) du dv,$$

where,

$$\Upsilon_{E,1}(u, v) = -\frac{1}{\pi^2(u-v)^2} + \frac{3}{2\pi^4(u-v)^4} + \mathcal{E}(u, v) \\ + \frac{1}{N^2\kappa_E^2} \left(F_1(u, v) + F_2(u, v) \sum_{i,j} \mathcal{C}_4(H_{ij}) + F_3(u, v) \sum_i \mathcal{C}_3(H_{ii}) \right)$$

and

$$\Upsilon_{E,2}(u, v) = -\frac{1}{2\pi^2(u-v)^2} + \mathcal{E}(u, v) \\ + \frac{1}{N^2\kappa_E^2} \left(\frac{1}{2}F_1(u, v) + F_2(u, v) \sum_{i,j} \mathcal{C}_{2,2}(H_{ij}) + F_3(u, v) \sum_i \mathcal{C}_3(H_{ii}) \right).$$

Here $\mathcal{C}(\cdot)$ is the cumulant, $\kappa_E := 2\pi\varrho_E = \sqrt{4 - E^2}$, and F_1, F_2, F_3 are explicit bounded elementary functions.

Blue = Average of Y_β , Red = non-universal, $\mathcal{E}(u, v)$ = quantitatively controlled error term.

Remarks

- **WGDM statistics** valid to leading order on all mesoscopic scales $N\omega \ll N$. Fails at macroscopic scale $N\omega \asymp N$.
- For $\beta = 1$, the **subleading corrections of WGDM** are leading order for $N\omega \ll \sqrt{N}$. For $N\omega \gg \sqrt{N}$, dominant subleading corrections are **non-universal**. These become leading order on macroscopic scales $N\omega \asymp N$.
- Result completely insensitive to size of spectral window $N\eta \gg 1$. Even at macroscopic scale $N\omega \asymp 1$, our result is much more precise than [Lytova, Pastur, 2009].
- After the local rescaling by $N\rho_E$, the **WGDM terms** do not depend on E but the **non-universal** ones are proportional to ρ_E^{-2} .

Comparison to Gustavsson's theorem

Gustavsson [2005] analyses mesoscopic correlations of eigenvalue **locations** instead of **densities** for GUE.

Order eigenvalues $\lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_N$ and introduce quantiles γ_i defined by $i/N = \int_{-\infty}^{\gamma_i} \rho_x dx$. Define normalized eigenvalues

$$\tilde{\lambda}_i := \frac{\pi \rho_{\gamma_i} N (\lambda_i - \gamma_i)}{\sqrt{\log N}}.$$

Theorem. $\tilde{\lambda}_i \sim \mathcal{N}(0, 1)$ and $\text{Cov}(\tilde{\lambda}_i, \tilde{\lambda}_j) \sim \beta$, where $\beta = 1 - \log_N(j - i)$.

Locations of two mesoscopically separated eigenvalues have a covariance of the same order as their individual variances.

Interpretation: eigenvalues fluctuate as a semi-rigid jelly on the scale $\sqrt{\log N}/N$.

Fluctuations of eigenvalue locations and density have little to do with one another.

Some ideas of proof

We can rewrite

$$p_E(u, v) = \frac{1}{(N \varrho_E)^2} \text{Cov}(X^\eta(u), X^\eta(v)),$$

where $X^\eta(u) := \sum_i f^{\eta/\varrho_E} \left(E + \frac{u}{N \varrho_E} - \lambda_i \right)$ and $f^\varepsilon(x) := \frac{1}{\varepsilon} f\left(\frac{x}{\varepsilon}\right)$.

Key difficulty, appearing throughout the proof: we are computing the covariance (to an arbitrary precision) of two weakly correlated random variables:

$$\text{Cov}(X^\eta(u), X^\eta(v)) \asymp \frac{1}{\omega^2}, \quad \text{Var}(X^\eta(u)) \asymp \frac{1}{\eta^2},$$

with $1/N \ll \eta \leq \omega \leq 1$.

Main work: compute covariance of Green functions (with $E_1 - E_2 = 2\omega$)

$$G = (H - E_1 - i\eta)^{-1}, \quad F = (H - E_2 - i\eta)^{-1}.$$

With notations

$$\underline{M} := \frac{1}{N} \operatorname{Tr} M, \quad \langle X \rangle := X - \mathbb{E}X,$$

we have to compute

$$\mathbb{E}\langle \underline{G} \rangle \langle \underline{F}^* \rangle.$$

We do this by deriving a **recursive family of self-consistent equations**, indexed by a finite tree, in polynomials in expectations of polynomials of the variables \underline{A}^m , $\langle \underline{A}^m \rangle$, A_{xy}^m , where $A = G, G^*, F, F^*$.

Simple tools

- Resolvent identity $zG = H(H - z)^{-1} - I$.
- Cumulant expansion [Khorunzhy, Khoruzhenko, Pastur – 1996]

$$\mathbb{E}[h \cdot f(h)] = \sum_{k=0}^l \frac{1}{k!} \mathcal{C}_{k+1}(h) \mathbb{E}[f^{(k)}(h)] + (\text{Error}).$$

Can be viewed as a generalization of Gaussian integration by parts to arbitrary random variables. Alternatively, a quantitative and more precise version of Stein's method.

Very powerful for deriving recursive high moment estimates in RMT [He, K – 2016]. This strategy was subsequently used to derive local laws [Lee, Schnelli – 2017], [He, K, Rosenthal – 2017].

How to start

Apply resolvent identity and cumulant expansion to $\mathbb{E}\langle \underline{G} \rangle \langle \underline{F}^* \rangle$, and get

$$\mathbb{E}\langle \underline{G} \rangle \langle \underline{F}^* \rangle = \frac{1}{-z_1 - 2\mathbb{E}\underline{G}} \left(\frac{2}{N^2} \mathbb{E}\underline{G}\underline{F}^{*2} + \frac{1}{N} \mathbb{E}\langle \underline{G}^2 \rangle \langle \underline{F}^* \rangle + \mathbb{E}\langle \underline{G} \rangle^2 \langle \underline{F}^* \rangle + \mathcal{W}_1 \right).$$

Leading term of order $1/(N\omega)^2$ (will have to be computed precisely by another self-consistent equation).

Error terms have to be estimated. Naive attempt: using the local semicircle law, we obtain

$$\mathbb{E}\langle \underline{G} \rangle^2 \langle \underline{F}^* \rangle = O\left(\frac{1}{(N\eta)^3}\right).$$

Much too big!

Reason: we did not exploit that $\langle \underline{G} \rangle^2$ and $\langle \underline{F}^* \rangle$ are weakly correlated. Solution: self-consistent equations for error terms.

General scheme for a term X :

- If X is an error term and the naive bound is too large, derive a self-consistent equation that expresses X in terms of a family of other terms.
- If X is a term we wish to compute, derive a self-consistent equation that extracts its main contribution plus error terms.

At each step, every term X gives rise to a set of children $\mathcal{S}(X)$ of further terms (tree).

How do we stop?

1. Identify a large enough set \mathcal{F} of terms that is closed under the map $X \mapsto \mathcal{S}(X)$.
2. Find bounds that allows to estimate all $X \in \mathcal{F}$ of a sufficiently high generation.

Algebra gets somewhat involved.

Tools for stopping

For the stopping in 2, we need much more than the local semicircle law: a priori bounds from [He, K – 2016] on $(G^m)_{ij}$, and the estimate

$$\mathbb{E}\underline{G}^m = O_m(1) \quad (2)$$

for all $m \in \mathbb{N}$.

(Local semicircle law gives $\mathbb{E}\underline{G}^m = O(1/\eta^{m-1})$.)

Interpretation: expected eigenvalue density ρ has **uniformly bounded derivatives of all order** down to all mesoscopic scales:

$$\begin{aligned} \operatorname{Im} \mathbb{E}\underline{G}^m &= \operatorname{Im} \int \frac{\rho(x)}{(x-z)^m} dx = (-1)^{m-1} \operatorname{Im} \int \frac{\rho_1^{(m-1)}(x)}{x-z} dx \\ &= (-1)^{m-1} \int \rho^{(m-1)}(x) \cdot \frac{\eta}{(x-E)^2 + \eta^2} dx = (-1)^{m-1} (\rho^{(m-1)} * \delta^\eta)(E). \end{aligned}$$

Proof: another recursive family of self-consistent equations, except that now it is possible to stop just using the local semicircle law and [He, K – 2016].

Define

$$F_i(u, v) = g_i\left(\frac{u - E}{N\rho_E}, \frac{v - E}{N\rho_E}\right).$$

where

$$g_1(x_1, x_2) = -\frac{4(4 + x_1x_2 + \sqrt{(4 - x_1^2)(4 - x_2^2)})}{\sqrt{(4 - x_1^2)(4 - x_2^2)}(\sqrt{4 - x_1^2} + \sqrt{4 - x_2^2})^2},$$

$$g_2(x_1, x_2) = \frac{2(x_1^2 - 2)(x_2^2 - 2)}{\sqrt{(4 - x_1^2)(4 - x_2^2)}},$$

$$g_3(x_1, x_2) = \frac{x_1^2x_2 + x_1x_2^2 - 2x_1 - 2x_2}{\sqrt{(4 - x_1^2)(4 - x_2^2)}}.$$

Error term

$$\mathcal{E}(u, v) = O\left(\frac{1}{(u - v)^5} + \frac{1}{N(u - v)^3} + \frac{1}{N^{3/2}(u - v)^2} + \frac{1}{N^2(u - v)}\right).$$